A Comparative Analysis of University Entrance Examinations Using the Construct Comparability Approach

Análisis comparativo de las Pruebas de Acceso a la Universidad bajo el enfoque de comparabilidad de constructo


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Abstract
Evaluation processes are a fundamental tool for training and selecting students. However, there are no empirical studies in Spain that analyse the usefulness of assessment tests to measure academic performance. This study, based on research carried out using the construct comparability approach, conducts a comparative analysis of grades achieved by 6,709 students pertaining to 15 academic subject areas of the university entrance examinations (Pruebas de Acceso a la Universidad - PAU) administered in the province of Alicante (Spain). The partial-credit Rasch model is used as an estimation method in which each academic subject is regarded as an instrument item related to the measurement of the academic performance construct. The initial results exhibited unidimensionality, and all academic subject areas fit the model, although there was a lack of discrimination between high- and low-performing students, mainly due to the absence of monotonicity in the scoring categories. The difficulty levels

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of the academic subjects were found to be appropriate for the skill levels of most students. These results demonstrated the ability of the tests analysed to report on the academic performance of the students who took the tests. In addition, important conclusions are presented regarding improvements in the grading processes, and future research studies are proposed.

**Keywords:** university entrance examination, educational assessment, academic performance, construct comparability approach, partial-credit Rasch model.

**Introduction**

In recent years, there has been an ongoing boom in the study of academic performance at all levels of education. Some studies analyse the cognitive, motivational, and contextual variables involved at the
predictive or causal levels (Dicke et al., 2018; Valle et al., 2008), while others analyse the quality of the measurement of external performance assessment tests and their associated variables (Martí and Puertas, 2018; Sayans-Jiménez, Vázquez-Cano, and Bernal-Bravo, 2018). In the latter case, it is worth highlighting the progress in research on the design and implementation of internationally standardised tests, such as the TIMMS (Trends in International Mathematics and Science Study), the PIRLS (Progress in International Reading Literacy Study), the IALS (International Assessment of Literacy Survey), and, especially, the PISA (Programme for International Student Assessment). However, it should be noted that in Spain, despite in-depth studies of various variables used in the analysis of the abovementioned tests (Elosua, 2013), there has been scant analysis in recent years of the processes for measuring the quality of the PAU (Pruebas de Acceso a la Universidad) university entrance examinations beyond some quantitative analyses of the differences between groups of test subjects in specific areas, or of a local nature (Rodríguez-Menéndez, Inda and Peña-Calvo, 2014; Ruiz et al., 2011).

The PAU examination is the current procedure for students who have earned a Spanish baccalaureate to access university studies in Spain and its territories. This examination supports the university admission of students through a series of core and elective academic subject tests. These tests have different formats depending on the academic subject at hand, including text or image commentaries, (short or long) essays on a specific subject, or problem solving, among others. In addition, these test designs are independent of the country’s various Autonomous Communities (students from each Community take the same tests), and they produce grades that, when weighted with upper secondary school transcript grades, are used to calculate a total grade used in the students’ applications for admission to the various university degree programmes. This process is based on the regulation established in Organic Law 8/2013, of 9 December, for the improvement of educational quality (LOMCE, 2013), and Royal Decree 412/2014, of 8 June, which establishes the basic regulations for admission to official university degree programmes.

Since the PAU is a key assessment process for the futures of thousands of students, it is imperative to consider the role of evaluative research in the field of education. In this sense, it is necessary to begin with a pragmatic and contextual approach for examining the processes and the results obtained, as well as their use in the methods used by various
organisations (Sondergeld and Koskey, 2011) to ensure the principles of equity and equal opportunity for university admission.

In the quantitative research field, various statistical methods have been applied to investigate the fulfilment of the conditions necessary to ensure an objective measurement of academic performance, as well as the correct design of measurement instruments based on the analysis of the most specific conditions. Noteworthy are the use of added-value models and multilevel models for the analysis of longitudinally measured academic performance (Blanco, González, and Ordóñez, 2009; López-Martín, Kouosmanen, and Gaviria, 2014). More specific to the PAU examination is the notable research by Gaviria (2005), in which he uses various statistical techniques (classical method, ordinary least squares method, multilevel method, average equalisation method, and standard deviation) to analyse the equivalence of baccalaureate grades with PAU exam grades, the latter serving as an anchoring point since it is a standardised test for all students. The results show that non-classical methods produce better results than the classical weighting method and make for a fairer student selection process.

However, while the PAU is useful as a standard set of tests for all students (taking into account the necessary knowledge domain-related differences), it is important to ensure that the tests adequately measure students’ various ability levels and that they demonstrate an adequate distribution of difficulty levels.

The scientific literature in the academic certification test field includes relevant research conducted in other countries that attempts to analyse the previously described psychometric properties based on different theoretical models for analysing the comparability of academic results. Particularly noteworthy are the models developed in the United Kingdom; more specifically, the performance comparability approach (Baird, Cresswell, and Newton, 2000), the conventional or sociological comparability approach (William, 1996b), the statistical comparability approach (William, 1996a), and a more recently developed model that improves on the previous ones – the construct comparability approach (Newton, 2005). This last model requires that comparisons of two elements have something in common that serves as the basis for this comparison. Just as two tests can be compared by measuring them on the same scale, in the context of academic score comparisons, we can only compare those that measure a common construct – academic performance, in our case.
Thus, the premise of this approach is as follows (Coe, 2008): two scores pertaining to two students are comparable if the academic performance of both (which corresponds to the same level of the latent construct they share) is measured using the same scoring method. According to this postulation, subject matter difficulty will correspond to a specific level established in the latent variable. That is, one academic subject is considered more difficult than another if, to achieve a specific score, a higher level of performance or ability is necessary (Coe, 2010).

Other studies have also demonstrated the need to analyse the usefulness of academic certification tests designed for student admission and to ensure comparability of results. For example, Hübner, Wagner, Hochweber, Neumann, and Nagengast (2019) showed that the results from two tests taken by students in Germany could lead to a deficient admission process because these tests had not been revised to reflect the educational reforms implemented. Furthermore, Korobko, Glas, Bosker, and Luyten (2008) found that the results achieved by students from the Netherlands were influenced by the academic subjects chosen for assessment, revealing the need to adjust the evaluation procedure to avoid an unfair assessment of higher-performing students.

Thus, considering the relevant research on this topic, a comparability measure must use the scores from academic subject tests as an instrument to measure the validity of the construct. This necessity implies that they must adequately represent the content, have good internal consistency, and demonstrate an appropriate degree of correlation between the latent construct and the variables (the different academic subjects).

This measurement model would be impossible without a clear conceptualisation of the construct – academic performance, in our case. It is important to note that, despite being a widely studied concept, there is no single definition of academic performance in the scientific literature. Given the complexity and multidisciplinary approach of academic performance, most of the working definitions refer to the assessment or evaluation of overall achievement at the primary-school level (Jiménez, 2000). However, our construct refers to the achievement levels attained as measured by the respective evaluation standards for the various academic subjects that comprise the PAU tests. This achievement level is translated into specific scores such that a comparison of the construct can be made if increasing or decreasing a score also means progressing or regressing in the construct that is measured.
The theoretical approach presented in this study adheres to the measurement approach espoused by the Rasch model (Rasch, 1980; Wright and Stone, 1979), which is the best-known example of item response theories (IRT). The Rasch model provides a mathematical model based on the calibration of ordinal data from a common measurement scale and checks for conditions such as unidimensionality, linearity, and monotonicity. In its most basic form, this model establishes that the difficulty of an item and the ability of a human subject can be measured on a common scale and that the probability of a person answering correctly will be conditioned by the difference between his or her ability and the item difficulty. Both measures (ability and difficulty) are examined in logit units, since the model uses a logarithmic scale. The use of a common measurement scale allows the setting of homogeneous intervals, such that the difference between the parameters of item difficulty and student ability indicates the same probability of success throughout the whole scale.

At this level of analysis, the starting point is to treat each academic subject as a specific item with 0-10 scoring intervals that indicate different degrees or categories of success. The partial-credit model (Wright and Masters, 1982) enables us to individually analyse the difficulty of attaining a specific score for each academic subject, per the Rasch methodology. This methodology has been used in the UK to analyse the comparability of the General Certificate of Secondary Education certification tests for 16-year-olds and the General Certificate of Education-Advanced tests for 18-year-olds (Coe, 2008; He, Stockford, and Meadows, 2018). The model formula is as follows:

\[
\ln \left( \frac{P_{nij}}{P_{ni(j-1)}} \right) = B_n - D_i F_{ij} = B_n - D_{ij}
\]

where:
- \(P_{nij}\) is the probability that subject \(n\) will correctly answer item \(i\) in category \(j\);
- \(B_n\) is the measured ability of subject \(n\);
- \(D_i\) is the measured difficulty of item \(i\); and
- \(F_{ij}\) is the adjustment measured for item \(i\) in category \(j\) relative to category \(j-1\), which is the point at which categories \(j-1\) and \(j\) are equally likely in relation to the item measurement (Bond and Fox, 2007).
Thus, our main task in this study is the application of the construct comparability approach, developed over the last few decades in the United Kingdom, to the PAU tests administered in one of Spain’s Autonomous Communities. Specifically, the objectives are as follows: 1) to compare the adjustment levels and difficulty parameters of the various subjects and 2) to compare the distribution of difficulty levels to the scores achieved in the different academic subjects across the latent attribute.

Methodology

Sample

The sample consists of most of the students from Alicante Province who took the PAU tests in June 2018. Specifically, test scores were collected from 6,709 students who were tested in the province’s two public universities: The University of Alicante and Miguel Hernandez University of Elche. The percentage of women in both universities is approximately 60%. The test scores were obtained from the university regulatory agency of the Valencia regional government.

Instruments

The study used the PAU tests administered in Alicante Province in June 2018. These are the same tests administered in the Autonomous Community’s other provinces (Valencia and Castellón). Fifteen academic subjects were selected from the set of core and elective subject tests administered. The selection criterion was a minimum of 600 students tested per academic subject. This criterion is used to ensure greater precision of the estimated parameters (He, Stockford, and Meadows, 2018). Thus, the academic subjects selected were Biology, Castilian Language and Literature, Audio-Visual Culture II, Technical Drawing, Business Economics, Physics, Geography, Art History, History of Spain, History of Philosophy, English, Latin II, Mathematics II, Applied Mathematics for the Social Sciences II, Chemistry, and Valencian Language and Literature.
The tests are scored based on standards previously established by each grading commission. The scoring criteria stem from the maximum possible score for each test question, together with qualitative instructions to reinforce the objectivity of the examiners. These scoring criteria are public and available from the Generalitat Valenciana website (http://www.ceice.gva.es/va/web/universidad/examenes-y-criterios-de-correccion-de-convocatoria-ordinaria)

**Procedure**

The construct comparability approach was applied in this study, which assumes that the grades achieved by students in different academic subjects as part of the overall testing process can be compared with one another.

The partial-credit model was applied using the Winsteps (version 4.4.0) statistical software package (Linacre, 2019) for the joint maximum likelihood estimation (Bond, 2004). In this model, each of the included academic subjects is regarded as an item of a single instrument that is capable of measuring the academic performance construct.

First, per the Rasch model, the model’s unidimensionality is measured by analysing the main components of the residual scores. According to Linacre (1998), the eigenvalue obtained by comparing the residuals should not be higher than 2.

The process of estimating the item difficulty (including their respective categories) and student ability levels is iterative and examines the relationship between the probabilities of obtaining a certain score as a function of the student’s ability. The maximum likelihood procedure enables us to determine the difficulty value of a certain score that best explains the recorded pattern of performance. Similarly, the ability value can be determined for each individual based on the pattern of difficulty indices. This process is repeated using the ability and difficulty estimates until they converge.

While many statistical models try to fit the model to the data, the opposite occurs in this model. In other words, the data must fit the model to be accepted. This fit can be determined from the residual measurements; that is, from the difference between a student’s answer for a given item and the expected answer calculated by the model. There
are two ways to standardise the fit measurements for a particular item or subject (Bond and Fox, 2007):

- **Outfit** is the quadratic average of the residuals, divided by the degrees of freedom. This statistic can be interpreted as an overall measure of whether the answers given to a particular item fit the model.

- **Infit** eliminates the extreme scores that influence the outfit, thereby utilising the residuals of the individuals whose ability levels are in the range closest to the specific item.

The infit and outfit statistics are calculated using quadratic averages as a function of the Pearson’s Chi-squared statistical value divided by its degrees of freedom, thus producing a scale with values ranging from 0 to infinity. Values below 1 indicate a better-than-expected fit to the model, while values above 1 indicate a poor fit to the model. Thus, if we have an infit value of 1.40, we can assert that there is 40% more data variability compared to the model prediction, while an outfit of 0.80 indicates that there is 20% less data variability with respect to the model prediction.

Different fit values have been established depending on the purpose of the analysis (Coe et al., 2008; Tan and Yates, 2007). Linacre (2002) suggested that values greater than 2 necessarily indicate a poor model fit, and therefore the analysis does not provide reliable conclusions. For this reason, we used this parametric adjustment value for the test items and students, which aligns with previous studies that have also applied the construct comparability approach (He, Stockford, and Meadows, 2018). In addition, the students’ average ability level for the various academic subjects was set to zero so that the parameter estimates could be compared with one another.

**Results**

First, the overall statistics for the model showed a subject (student) reliability index of 0.74 and a subject (student) separation index of 1.69. These values are considered low and indicate that the set of academic subjects is not sensitive enough to effectively distinguish between high- and low-performing students (Bond and Fox, 2007).
Regarding the model’s unidimensionality, the results of the analysis of the main components of the residual scores (Bond and Fox, 2007) reveal a main factor that can explain 51.3% of the variance in the latent attribute. The value of the hypothetical second factor is lower than 2 (Eigenvalue V2 = 1.4), which confirms the model’s unidimensionality.

Table I presents the academic subjects in order of difficulty (from highest to lowest), as well as their respective goodness-of-fit indices. An optimal fit of all academic subjects to the model is observed in accordance with the established criteria. The academic subjects with the highest difficulty indices are (in descending order) Chemistry, Geography, and Physics. The academic subjects with the lowest difficulty indices are (in ascending order) History of Spain, Mathematics II, and Economics.

<table>
<thead>
<tr>
<th>Academic subjects</th>
<th>No. of students</th>
<th>Difficulty</th>
<th>Infit</th>
<th>Outfit</th>
<th>Item-scale correlation</th>
</tr>
</thead>
<tbody>
<tr>
<td>Chemistry</td>
<td>2,179</td>
<td>-0.03</td>
<td>0.92</td>
<td>0.91</td>
<td>0.76</td>
</tr>
<tr>
<td>Geography</td>
<td>1,619</td>
<td>-0.16</td>
<td>1.40</td>
<td>1.40</td>
<td>0.62</td>
</tr>
<tr>
<td>Physics</td>
<td>1,504</td>
<td>-0.24</td>
<td>1.12</td>
<td>1.12</td>
<td>0.75</td>
</tr>
<tr>
<td>Applied Mathematics for the Social Sciences II</td>
<td>2,236</td>
<td>-0.26</td>
<td>1.25</td>
<td>1.25</td>
<td>0.63</td>
</tr>
<tr>
<td>Biology</td>
<td>1,738</td>
<td>-0.29</td>
<td>0.90</td>
<td>0.89</td>
<td>0.74</td>
</tr>
<tr>
<td>Technical Drawing</td>
<td>697</td>
<td>-0.31</td>
<td>1.51</td>
<td>1.50</td>
<td>0.62</td>
</tr>
<tr>
<td>Art History</td>
<td>714</td>
<td>-0.41</td>
<td>1.30</td>
<td>1.29</td>
<td>0.68</td>
</tr>
<tr>
<td>Castilian Language and Literature II</td>
<td>6,123</td>
<td>-0.52</td>
<td>0.71</td>
<td>0.74</td>
<td>0.67</td>
</tr>
<tr>
<td>Valencian Language and Literature II</td>
<td>4,747</td>
<td>-0.55</td>
<td>0.69</td>
<td>0.71</td>
<td>0.64</td>
</tr>
<tr>
<td>Latin</td>
<td>864</td>
<td>-0.63</td>
<td>1.17</td>
<td>1.16</td>
<td>0.69</td>
</tr>
<tr>
<td>English</td>
<td>5,960</td>
<td>-0.65</td>
<td>1.17</td>
<td>1.16</td>
<td>0.61</td>
</tr>
<tr>
<td>History of Philosophy</td>
<td>778</td>
<td>-0.66</td>
<td>1.37</td>
<td>1.34</td>
<td>0.64</td>
</tr>
<tr>
<td>Economics</td>
<td>1,698</td>
<td>-0.67</td>
<td>0.99</td>
<td>0.98</td>
<td>0.70</td>
</tr>
<tr>
<td>Mathematics II</td>
<td>3,177</td>
<td>-0.75</td>
<td>1.24</td>
<td>1.19</td>
<td>0.68</td>
</tr>
<tr>
<td>History of Spain</td>
<td>6,124</td>
<td>-0.81</td>
<td>0.90</td>
<td>0.92</td>
<td>0.62</td>
</tr>
</tbody>
</table>

Source: Created by the authors based on results obtained using the Winsteps software package.
Since the partial-credit model was used, all of the academic subject scores have their own goodness-of-fit indices. In this sense, all of the academic subject scores have an optimal fit, with values ranging from 0.7 to 1.9. However, the score difficulty distribution presented in Figure I shows that the monotonicity criterion was not met for several academic subjects. Thus, no increase in the associated difficulty was detected when moving from one score to the score immediately above it. This situation occurred for the academic subjects of the History of Spain, History of Philosophy, Technical Drawing, and Geography. It was also observed that the scores at the top and bottom of the scale exhibit greater dispersion in the scalar distribution. For example, achieving a score of 10 in Biology is more difficult than achieving a score of 10 in Applied Mathematics for the Social Sciences. Similarly, achieving a score of one in Latin is more difficult than achieving the same score in Valencian Language and Literature II.

![Figure I. Test score distribution for latent attributes by academic subject](image-url)
The “Wright map” presented in Figure II shows the distribution of persons and items along the ranges of ability and difficulty, respectively. Individuals are distributed on the left side of the graph, while items are on the right side. The figure shows that the difficulty of the different academic subjects corresponds to an ability range between logits 0 and -1. This result is positive, since it means that most students are able to take all the exams. There is also a percentage of students between logits 0 and 1, which indicates a higher level of ability and therefore a higher probability of achieving good test scores.

FIGURE II. Person-Item map

Source: Created by the authors based on results obtained using the Winsteps software package

Lastly, Figure III shows the item characteristic curves (ICC) for all of the academic subjects analysed. These curves reveal the relationships between the expected and observed scores for each item, depending
on the $q$ student ability level in the latent attribute. There is good fit when the observed and expected scores overlap between the points and the line, respectively. Specifically, a similar academic subject difficulty pattern is identified in the average scores (between -2 and +2 logits) and in the differences in the maximum and minimum scores (between -4 and -2 logits and between +2 and +2 logits). The easiest academic subjects (based on required ability level) are displayed on the left side of the graph, while the scores that require a higher ability level are on the right.

**FIGURE III. Item Characteristic Curves**

![Item Characteristic Curves](image-url)
Discussion and conclusions

The initial analyses demonstrate compliance with the criterion of unidimensionality, which is essential for the application of the model and for establishing a latent construct defined as academic performance within the scope of the PAUs. However, despite the creation of this operational construct, it should not be assumed that a single, overall process exists. The scientific literature indicates that the interpretation of this construct is not clear, since it is not a basis for establishing the specific purposes of each test developed (He, Stockford and Meadows, 2018). Therefore, it is clear that all tests require specific skills, but they all also require global cognitive processes related to the measurement of the construct.

Regarding the first objective, we observed an optimal fit of all the academic subjects analysed, which enables an examination of the invariance properties assumed by the Rasch model (Bond and Fox, 2007). Therefore, the value of this type of estimation lies in the potential for making inferences beyond the sample of students used. At the same time, the academic subject fit allows comparisons with the difficulty parameters obtained from the ability level required to achieve a certain test score. These results lead to a key conclusion in the field of PAU assessment regarding the choice of academic subjects by students, a point that has been widely discussed in the international body of literature (Lamprianou, 2009). Bell et al. (2007) indicate that a student’s perceived difficulty of one or more academic subjects may pose a barrier to university admission. Consequently, other academic subjects have higher enrolment rates. Considering the results of this study, such a situation could be occurring with the History of Spain, to the detriment of the History of Philosophy, since students must choose one of these academic subjects, and the former has more than three times as many candidates as the latter.

The analysis of the second objective highlights the need to consider the traditional grading scale used in the PAUs, as the typical monotonicity criterion was not met in the performance tests. The current 10 categories do not adequately discriminate at specific points of the latent attributes. It should be noted that in most countries where comparative analyses of admission test results have been conducted, fewer grade categories are used. Adjusting the scoring system would allow other possible
calculations in the same comparative framework, such as evaluating the relative difficulty of each score by comparing it with the average difficulty of all academic subjects; these differences could be expressed in the logit unit of measurement or in terms of direct scores. This measure would enable observations of how the difficulty of the academic subject scores evolve over successive testing sessions in different academic years (He, Stockford and Meadows, 2018).

Regarding the previous paragraph, the separation rate obtained is low, which affects the fact that the evidence does not discriminate well between students who demonstrate high and low levels of the latent attribute. However, the “Wright map” reveals that all the test difficulty levels are within the students’ ability range, so there are adequate degrees of probability of obtaining positive results. The positions of the academic subjects on the scale correspond to a similar distribution of the categories in the latent construct (as seen in the ICC), although some differences can be noted in the distribution of the extremity categories. Once again, these results highlight the need to recodify the categories and improve discrimination by including more students in each high- and low-performance category.

In conclusion, this study aims to initiate an effective analysis in Spain that compares test scores using the construct comparability approach that has been applied in other countries. However, it is important to bear in mind certain limitations that may guide future research on this topic. First, it should be noted that the national data samples used in other countries are much larger. Larger samples enable better estimates because of the higher number of test scores and academic subjects. Conducting this study at a provincial level allowed the use of a construct comparability approach and confirmed the possibility of conducting future studies analogous to those conducted in other countries. In our specific context, this approach enables comparisons between Autonomous Communities to determine appropriate measures for equity. In this sense, the influences of several differential factors must be analysed, such as the individual selection of academic subjects or the effects of educational reforms on testing (Hübner et al., 2019; Korobko, Glas, and Bosker, 2008). Nonetheless, the potential influence of the test graders on the model measurement was not examined in this study and has not yet been explored in the field’s scientific literature. However, considering that most PAUs consist of written essay questions, the differences between
test graders regarding task interpretation and evaluation categories (as well as other possible effects such as the halo effect, gender and cultural bias, etc.) can contribute to measurement error and to the validity and fairness of test scoring (Prieto, 2011). The Rasch model can incorporate these types of considerations through an extension of the partial-credit model called the Many-Facet Rasch Measurement.

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